POSTAL Book Package

2023

ESE

Electronics Engineering

Conventional Practice Sets

Communication Systems

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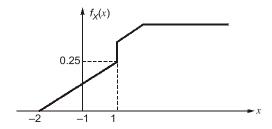


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Theory of Random Variable and Noise

Q.1 Define PDF and summarise its important properties. Also calculate the probability of outcome of a Random Variable (RV) X having $X \le 1$ for the following PDF curve of RV as shown.



Solution:

Probability density function specifies the probability of a random variable taking a particular value.

The Probability Density Function (PDF) which is generally denoted by $f_{\chi}(x)$ or $P_{\chi}(x)$ or $P_{\chi}(x)$ is defined in terms of the Cumulative Distribution Function (CDF) $F_{\chi}(x)$ as,

$$PDF = f_X(x) = \frac{d}{dx} F_X(x)$$
...(i)

The PDF has the following properties:

- (i) $f_x(x) \ge 0$ for all x
 - This results from the fact that probability cannot be negative. Also, $F_{\chi}(x)$ increases monotonically, as x increases, more outcomes are included in the prob. of occurrence represented by $F_{\chi}(x)$.
- (ii) Area under the PDF curve is always equal to unity.

i.e.
$$\int_{-\infty}^{\infty} f_X(x) \, dx = 1$$

(iii) The CDF is obtained by the result

CDF =
$$\int_{-\infty}^{x} f_X(x) dx$$

(iv) Probability of occurrence of the value of random variable between the limits of x_1 and x_2 is given by,

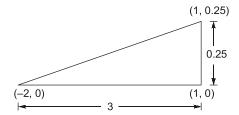
$$P(x_1 < X < x_2) = \int_{x_1}^{x_2} f_X(x) dx$$

Now consider the given PDF curve, since we have to find $P(x \le 1)$ so,

Equation for the PDF curve for $x \le 1$ is,

$$f_X(x) = \left(\frac{1}{12}x + \frac{1}{6}\right)$$





Now, $P(x \le 1)$

$$= P(-2 < x < 1) = \int_{-2}^{1} \left(\frac{1}{12}x + \frac{1}{6}\right) dx = \left[\frac{1}{12} \cdot \frac{x^{2}}{2} + \frac{1}{6}x\right]_{-2}^{1} = \frac{3}{8}$$

$$P(x \le 1) = \frac{3}{8}$$

Q2 Find the cumulative distribution function F(x) corresponding to the PDF $f(x) = \frac{1}{\pi(1+x^2)}, -\infty < x < \infty$.

Solution:

Given
$$f(x) = \frac{1}{\pi(1+x^2)}, -\infty < x < \infty$$

$$F(x) = P(X \le x)$$

$$= \int_{-\infty}^{x} f(x) dx = \frac{1}{\pi} \int_{-\infty}^{x} \frac{dx}{1+x^2} = \frac{1}{\pi} \left[\tan^{-1} x \right]_{-\infty}^{x} = \frac{1}{\pi} \left(\frac{\pi}{2} + \tan^{-1} x \right)$$

Q3 Given the random variable X with density function

$$f_{\chi}(x) = \begin{cases} 2x, & 0 < x < 1 \\ 0, & \text{otherwise} \end{cases}$$

Find the PDF of $Y = 8 X^3$.

Solution:

Given
$$y = 8x^{3} \text{ is an increasing function in } (0, 1)$$

$$y = 8x^{3}$$

$$\Rightarrow x^{3} = \frac{y}{8}$$

$$\Rightarrow x = \left(\frac{y}{8}\right)^{1/3} = \frac{1}{2}y^{1/3}$$
and
$$f_{\chi}(x) = 2x, \qquad 0 < x < 1$$

$$f_{\chi}(y) = \frac{2y^{1/3}}{2} = \frac{y^{1/3}}{3}$$

$$f_{\gamma}(y) = x = \left(\frac{y}{8}\right)^{1/3} = \frac{1}{2}y^{1/3} \qquad \Rightarrow \frac{dx}{dy} = \frac{1}{6}y^{-2/3}$$
Using it in (i)
$$f_{\gamma}(y) = y^{1/3} \frac{1}{6}y^{-2/3} = \frac{1}{6}y^{-1/3} = \frac{1}{6}\frac{1}{y^{1/3}} = \frac{1}{6}\frac{1}{3\sqrt{y}}$$

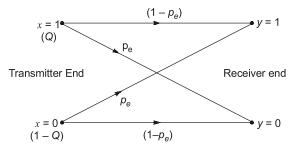
The range for x is 0 < x < 1



When x = 0, $y = 8 \times 0 = 0$ and x = 1, $y = 8 \times 1^3 = 8$

$$f_{\gamma}(y) = \frac{1}{6\sqrt[3]{y}}, \qquad 0 < y < 8$$

A BSC (Binary Symmetric Channel) error probability is P_e . The probability of transmitting '1' is Q, and that of transmitting '0' is (1 - Q) as in figure below. Calculate the probabilities of receiving 1 and 0 at the receiver?



Solution:

If x and y are the transmitted digit and the received digit respectively, then for a BSC,

$$P_{\mathcal{N}^{\chi}}(0|1) = P_{\mathcal{N}^{\chi}}(1|0) = P_{e}$$

$$P_{\mathcal{N}^{\chi}}(0|0) = P_{\mathcal{N}^{\chi}}(1|1) = 1 - P_{e}$$
Also,
$$P_{\chi}(1) = Q \text{ and } P_{\chi}(0) = 1 - Q$$
We have to find,
$$P_{\chi}(1) \text{ and } P_{\chi}(0) = ?$$

$$\therefore \qquad P_{\chi}(1) = P_{\chi}(0) P_{\mathcal{N}^{\chi}}(1|0) + P_{\chi}(1) P_{\mathcal{N}^{\chi}}(1|1) = (1 - Q)P_{e} + Q(1 - P_{e})$$
also,
$$P_{\chi}(0) = P_{\chi}(0)P_{\mathcal{N}^{\chi}}(0|0) + P_{\chi}(1)P_{\mathcal{N}^{\chi}}(0|1) = (1 - Q)(1 - P_{e}) + QP_{e}$$

Q5 For the triangular distribution

$$f(x) = \begin{cases} x, & 0 \le x \le 1 \\ 2 - x, & 1 \le x \le 2 \\ 0, & \text{otherwise} \end{cases}$$

Find the mean and variance.

Solution:

Mean =
$$E(X) = \int_{-\infty}^{\infty} xf(x)dx = \int_{0}^{1} x \cdot x \, dx + \int_{1}^{2} x(2-x) \, dx = \int_{0}^{1} x^{2} \, dx + \int_{1}^{2} (2x-x^{2}) \, dx$$

$$= \left[\frac{x^{3}}{3}\right]_{0}^{1} + \left[2\left(\frac{x^{2}}{2}\right) - \frac{x^{3}}{3}\right]_{1}^{2}$$

$$= \frac{1}{3} + \left[\left(4 - \frac{8}{3}\right) - \left(1 - \frac{1}{3}\right)\right] = \frac{1}{3} + \frac{4}{3} - \frac{2}{3} = 1$$

$$E(X^{2}) = \int_{-\infty}^{\infty} x^{2} f(x) dx = \int_{0}^{1} x^{2} x \, dx + \int_{1}^{2} x^{2} (2-x) \, dx$$

$$= \int_{0}^{1} x^{3} \, dx + \int_{1}^{2} (2x^{2} - x^{3}) \, dx = \left[\frac{x^{4}}{4}\right]_{0}^{1} + \left[2\left(\frac{x^{3}}{3}\right) - \frac{x^{4}}{4}\right]_{1}^{2}$$

$$= \frac{1}{4} + \left[\left(\frac{16}{3} - \frac{16}{4}\right) - \left(\frac{2}{3} - \frac{1}{4}\right)\right] = \frac{1}{4} + \frac{16}{3} - 4 - \frac{2}{3} + \frac{1}{4} = \frac{7}{6}$$

$$Var(X) = E(X^{2}) - E(X)^{2} = \frac{7}{6} - (1)^{2} = \frac{1}{6}$$



Q.6 The joint density function of two continuous random variables is given by

$$f(x, y) = \begin{cases} xy/8, & 0 < x < 2, \ 1 < y < 3 \\ 0, & \text{otherwise} \end{cases}$$

Find (a) E(X), (b) E(Y) and (c) E(2X + 2Y).

Solution:

(a)
$$E(X) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xf(x, y) dx dy = \int_{x=0}^{2} \int_{y=1}^{3} x(xy/8) dx dy = \frac{4}{3}$$

(b)
$$E(Y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} yf(x, y) dx dy = \int_{x=0}^{2} \int_{y=1}^{3} y(xy/8) dx dy = \frac{13}{6}$$

(c)
$$E(2X+3Y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (2x+3y) dx dy = \int_{x=0}^{2} \int_{y=1}^{3} (2x+3y)(xy/8) dx dy = \frac{55}{6}$$

Let z be a random variable with probability density function $f_z(z) = \frac{1}{2}$ in the range $-1 \le z \le 1$. Let the random variable x = z and the random variable $y = z^2$. Obviously x and y are not independent since $x^2 = y$. Show that x and y are uncorrelated.

Solution:

We have,
$$E(z) = \int_{-1}^{1} z \cdot f_{Z}(z) dz$$

 $\Rightarrow \qquad E(z) = \frac{1}{4} [z^{2}]_{-1}^{1} = 0$
Since, $x = z$, so $E(x) = E(z) = 0$
Since, $y = z^{2}$ so $E(y) = E(z^{2})$
So that, $E(y) = \int_{-1}^{1} \frac{1}{2} z^{2} dz = \frac{1}{6} [z^{3}]_{-1}^{1} = \frac{1}{3}$

We know that, the co-variance ' μ ' of two RVs x and y is defined as,

$$\mu = E\{(x - m_x) (y - m_y)\}\$$

$$= E\{(x)\left(y - \frac{1}{3}\right)\} = E\{xy - \frac{1}{3}x\} = E\{z^3 - \frac{z}{3}\} = \int_{-1}^{1} \frac{1}{2}\left(z^3 - \frac{z}{3}\right) dz$$

$$\mu = 0$$

Now, correlation coefficient between the variables x and y is defined by quantity ' ρ ' as,

$$\rho = \frac{\mu}{\sigma_r \sigma_v} = 0$$

So, we can say that these RV's X and Y are uncorrelated.

A WSS random process x(t) is applied to the input of an LTI system with impulse response $h(t) = 3e^{-2t} u(t)$

Find the mean value of the output y(t) of the system, if E[x(t)] = 2. Here $E[\cdot]$ denotes the expectation operator.



Solution:

The output y(t) is the convolution of the input x(t) and the impulse response h(t).

$$y(t) = \int_{-\infty}^{\infty} h(\tau) \cdot x(t - \tau) \cdot d\tau$$

$$E[y(t)] = \int_{-\infty}^{\infty} h(\tau) \cdot E[x(t - \tau)] \cdot d\tau$$

$$E[y(t)] = H(0) \times E[x(t)]$$

$$E[y(t)] = E[x(t)] \cdot H(0)$$

where, $H(0) = H(\omega)|_{\omega=0}$ and $H(\omega) =$ Fourier transform of h(t)

Given
$$E[x(t)] = 2$$
,

$$h(t) = 3e^{-2t}u(t)$$

Taking Fourier transform,

$$H(\omega) = \frac{3}{2 + i\omega} \implies H(0) = \frac{3}{2}$$

$$E[y(t)] = 2 \times \frac{3}{2} = 3$$

Suppose that two signals $s_1(t)$ and $s_2(t)$ are orthogonal over the interval (0, T). A sample function n(t) of a zero-mean white noise process is correlated with $s_1(t)$ and $s_2(t)$ separately, to yield the following variables:

$$n_1 = \int_0^T s_1(t) n(t) dt$$
 and $n_2 = \int_0^T s_2(t) n(t) dt$

Prove that n_1 and n_2 are orthogonal.

Solution:

$$E[n_{1}n_{2}] = E\left[\int_{0}^{T} s_{1}(u) n(u) du \int_{0}^{T} s_{2}(v) n(v) dv\right]$$
$$= \int_{0}^{T} \int_{0}^{T} s_{1}(u) s_{2}(v) E[n(u) n(v)] du dv$$

n(t) is a white noise process.

So,
$$R_{N}(\tau) = \frac{N_{0}}{2}\delta(\tau)$$

$$E[n(u)n(v)] = \frac{N_{0}}{2}\delta(u-v)$$
Hence,
$$E[n_{1}n_{2}] = \frac{N_{0}}{2}\int_{0}^{T}\int_{0}^{T}s_{1}(u)s_{2}(v)\delta(u-v)dudv$$

$$= \frac{N_{0}}{2}\int_{0}^{T}s_{1}(u)s_{2}(u)du$$

$$= 0 \qquad \therefore s_{1}(t) \text{ and } s_{2}(t) \text{ are orthogonal over } (0, T)$$

 $E[n_1n_2] = 0$. So, n_1 and n_2 are also orthogonal.